

The Stakeholder Illusion: Community Wind Ownership and the Limits of Financial Alignment

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Abstract

Can financial ownership of renewable energy infrastructure buy community acceptance? Denmark's *køberetsordning* (2009–2020) required wind developers to offer at least 20% of project shares at cost price to nearby residents, creating mandatory financial stakes in turbine returns. Using a staggered difference-in-differences design across 97 Danish municipalities—comparing 49 municipalities receiving new community-owned turbines to 48 controls—I find no significant effect on residential property values (Callaway-Sant'Anna ATT: -0.003 , 95% CI $[-0.019, 0.014]$) or green party vote shares (-0.62 pp, $p = 0.41$). A placebo test confirms that existing wind sites exhibit 10% lower property growth regardless of ownership structure. Financial co-ownership does not mitigate wind turbines' local disamenity costs or shift political preferences. The binding constraint on renewable energy acceptance appears to lie beyond financial incentives.

JEL Codes: Q42, Q58, R31, H23

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1. Introduction

The global energy transition requires building wind turbines where people live, but people who live near wind turbines oppose them. This tension—between aggregate climate benefits and local disamenity costs—is among the most consequential bottlenecks in decarbonization policy. In Denmark, the country that pioneered modern wind energy, approximately 30% of planned onshore projects faced significant local opposition during the 2010s (Borch, 2020), and the national parliament ultimately suspended the community ownership scheme in 2020 partly due to implementation frictions. The standard policy response is financial alignment: give neighbors a financial stake, and opposition should dissolve. But does it?

This paper tests whether mandatory community co-ownership of wind turbines reduces local opposition, as measured by property values and green party voting. I exploit Denmark’s *køberetsordning* (purchase-right scheme), which from 2009 to 2020 required developers of onshore turbines to offer at least 20% of project shares at cost price to residents within 4.5 kilometers. Using monthly wind capacity data from Energidataservice and property and election data from Statistics Denmark, I construct a panel of 97 municipalities observed from 2004 to 2024. Forty-nine municipalities received new community-owned wind capacity during 2017–2020, while 48 municipalities—including both historically wind-intensive and wind-free areas—serve as controls.

The main finding is a precisely estimated null. The Callaway-Sant’Anna aggregate average treatment effect on log property values is -0.003 (95% CI: $[-0.019, 0.014]$), ruling out effects larger than 2% in either direction. Green party vote share shows no significant response (-0.62 percentage points, $p = 0.41$). A naïve two-way fixed effects specification yields a significant negative coefficient (-3.8% , $p = 0.016$), but this reflects pre-existing differences in property value trajectories between wind-hosting and non-hosting municipalities rather than a causal effect of community ownership.

The placebo test is the sharpest diagnostic. Comparing municipalities with pre-existing (pre-2016) wind infrastructure to never-wind municipalities over the 2004–2016 period—before any new installations could confound the estimate—reveals that wind-hosting municipalities experienced property value growth approximately 10 percentage points lower than non-hosting municipalities. This effect is highly significant and of an order of magnitude larger than any community ownership effect. Wind turbines depress local property values regardless of who owns them.

This paper contributes to three literatures. First, it provides the first quasi-experimental evidence on whether financial community ownership reduces opposition to renewable energy infrastructure. The existing literature on community wind acceptance is entirely survey-based

(Jørgensen, 2020; Suskevics et al., 2019; Liebe et al., 2017; Wüstenhagen and Menichetti, 2012), making it impossible to distinguish stated from revealed preferences. By measuring property values and voting—both market-based outcomes that reflect actual behavior—this paper moves beyond hypothetical willingness-to-accept.

Second, it contributes to the literature on local opposition to infrastructure siting, building on Fischel (2001)’s “homevoter hypothesis” and the broader NIMBY literature (Dear and Wolch, 1992; Schively, 2007). The theoretical prediction is ambiguous: financial stakes could reduce opposition by aligning incentives (Coase, 1960; Dixit, 1996), but they could also be ineffective if opposition stems from non-financial concerns like landscape aesthetics, noise, or identity (Devine-Wright, 2009; Van der Horst, 2007).

Third, it contributes to the growing literature on property value impacts of wind turbines, joining Gielissen and Hoen (2024), Droller (2024), Hoen et al. (2015), and Sunakawa and Yamazaki (2023), among others. While these studies generally find negative property value effects near turbines, none isolates the ownership channel. My placebo test confirms the disamenity finding and shows that community ownership does not attenuate it. An important caveat: the municipality-level design identifies the net effect of the køberetsordning policy package (turbine installation with mandatory community ownership offer), not the pure ownership effect. The ideal within-project lottery comparison—winners versus losers among community share applicants—awaits individual-level data from Energistyrelsen’s lottery records.

The remainder of the paper proceeds as follows. Section 2 describes the institutional background of Denmark’s køberetsordning. Section 3 presents the data. Section 4 outlines the empirical strategy. Section 5 reports results. Section 6 discusses implications.

2. Institutional Background

Denmark’s wind energy landscape. Denmark has been a global leader in wind energy since the 1970s oil crisis, with onshore wind providing approximately 47% of electricity consumption by 2019 (Danish Energy Agency, 2020). The 2007 municipal reform consolidated Denmark’s 271 municipalities into 98, creating relatively large administrative units that form the basis of this analysis.

The køberetsordning (purchase-right scheme). Enacted as part of the 2008 Renewable Energy Act (VE-loven, §13–17), the køberetsordning required developers of all new onshore wind turbines above 25 meters to offer at least 20% of project shares at cost price to local residents. The pricing rule was based on the developer’s actual project costs, resulting in share

prices of approximately 3,000–4,000 DKK (roughly \$450–\$600) per share. Each participant could purchase up to 50 shares.

Allocation followed a geographic priority system. Residents within 4.5 kilometers of the turbine site (Group 1) had first priority. If shares remained, municipality residents outside the 4.5 km radius (Group 2) could purchase them, followed by residents of neighboring municipalities (Group 3). When demand exceeded supply—which occurred frequently, as the guaranteed feed-in tariff returns made shares attractive—Energistyrelsen (the Danish Energy Agency) conducted formal lotteries (lodtækning) to allocate shares.

The scheme covered approximately 90 projects between 2009 and 2020, distributing ownership stakes to thousands of households. It was suspended in June 2020, partly because administrative costs and legal disputes had increased, and partly because the shift to competitive auctions for wind capacity made the fixed-price share offering less practical.

The policy rationale. The køberetsordning was explicitly designed to reduce local opposition by giving residents a financial stake in turbine performance. The logic is straightforward: a household receiving annual dividend payments from a nearby turbine should weigh these returns against visual or noise disamenities, potentially shifting their net valuation from negative to positive. This “stakeholder alignment” hypothesis is the central premise of community benefit schemes worldwide, including Scotland’s community benefit register, Germany’s Bürgerbeteiligung models, and France’s investissement participatif requirements (Bauwens, 2016; Musall and Kuik, 2011; Walker, 2008).

3. Data

I combine three data sources to construct a municipality-year panel covering 2004–2024.

Wind capacity. Monthly onshore wind capacity data come from Energidataservice’s CapacityPerMunicipality dataset, which reports installed megawatts (MW) and turbine counts for each of Denmark’s 99 municipalities from January 2016 to February 2026. I use the January snapshot of each year as the annual observation. As of January 2020 (the last full year of the køberetsordning), 83 of 99 municipalities had some onshore wind capacity installed.

I define treatment based on changes in onshore wind capacity during the køberetsordning period. Municipalities are classified as “newly treated” if they received positive additions to onshore capacity between 2017 and 2020 (49 municipalities, total 664 MW of new capacity), “always treated” if they had existing capacity in 2016 but no new installations (35 municipalities), “never treated” if they had zero onshore wind throughout (14 municipalities), or “post-policy” if new capacity arrived only after 2020 (1 municipality, excluded).

Property values. Average assessed market values of single-family houses (enfamiliehuse) by municipality come from Statistics Denmark’s EJDFOE1 table, available annually from 2004 to 2024. This provides 21 years of outcome data, including 13 pre-treatment years for the earliest-treated municipalities.

Elections. Municipal election results by party come from Statistics Denmark’s VALGK3 table, covering five elections: 2005, 2009, 2013, 2017, and 2021. I construct a “green vote share” variable as the combined share of SF (Socialistisk Folkeparti), Alternativet, and Enhedslisten—De Rød-Grønne, the three parties most closely associated with climate and environmental policy in Danish politics.

Controls. Municipality-level population (FOLK1A, quarterly) and average personal income (INDKP106, annual) serve as time-varying controls.

Table 1 presents summary statistics for treated and control municipalities as of 2016. Treated municipalities have lower average property values (1,521,000 vs. 2,616,000 DKK), reflecting their more rural character, but larger populations on average. They also have substantially more existing wind capacity (62.3 vs. 15.8 MW), indicating that new installations tend to occur in municipalities already hosting wind infrastructure.

Table 1: Summary Statistics by Treatment Status (2016)

	Control (No new wind)	Treated (New wind 2017–2020)
Municipalities	48	49
Property value (1000 DKK)	2616	1521
SD	(1735)	(577)
Population	51167	65532
SD	(82889)	(54152)
Avg. income (1000 DKK)	235	214
Onshore wind capacity (MW)	15.8	62.3
N onshore turbines	26	97

Notes: Control municipalities include those with no wind (14) and those with existing wind but no new installations during 2017–2020 (34). Treated municipalities received new onshore wind capacity under Denmark’s køberetsordning (purchase-right scheme). Property values are average market valuations of single-family houses from DST EJDFOE1.

4. Empirical Strategy

4.1 Identification

I exploit the staggered timing of new onshore wind installations across municipalities during the k beretsordning period. The identifying assumption is that, absent new community-owned wind capacity, treated and control municipalities would have followed parallel property value and voting trajectories.

The primary estimator is Callaway and Sant’Anna’s (2021) group-time ATT, which avoids the forbidden-comparison bias inherent in two-way fixed effects (TWFE) estimation with heterogeneous treatment timing (Goodman-Bacon, 2021; de Chaisemartin and D’Haultf uille, 2020; Sun and Abraham, 2021). I define cohorts by the year of first new onshore wind installation (2017, 2018, 2019, or 2020). Not-yet-treated municipalities serve as the comparison group, which includes municipalities that will receive treatment later as well as municipalities that are never treated or already had wind before 2016.

4.2 Estimation

The baseline TWFE specification is:

$$\log(\text{PropertyValue}_{mt}) = \alpha_m + \lambda_t + \beta \cdot \text{Post}_{mt} + \varepsilon_{mt} \quad (1)$$

where Post_{mt} equals one if municipality m has received new onshore wind capacity by year t , α_m and λ_t are municipality and year fixed effects, and standard errors are clustered at the municipality level.

The Callaway-Sant’Anna estimator separately identifies $ATT(g, t)$ for each cohort g and time period t , then aggregates to obtain the overall ATT and dynamic (event-study) effects. I use outcome regression as the estimation method and compute simultaneous confidence bands via 1,000 bootstrap iterations.

4.3 Threats to validity

Two important clarifications about what this design identifies. First, the treatment variable captures the *joint* effect of receiving a new wind turbine with a mandatory community ownership offer. Since all new onshore turbines during 2009–2020 were subject to the k beretsordning, separating the disamenity effect of turbine siting from the compensating effect of ownership is not possible with municipality-level data alone. The ideal test—a within-project comparison of lottery winners and losers among applicants for community

shares—requires individual-level lottery records from Energistyrelsen, which are not publicly available. What this paper can test is the *net* effect: whether the k beretsordning policy package (turbine + mandatory ownership offer) generates different property value and voting outcomes than the counterfactual of no new turbine.

Second, the main threat is endogenous siting: municipalities receiving new wind capacity may differ systematically from controls. The pre-treatment balance (Table 1) confirms that treated municipalities are more rural and already have more wind infrastructure. I address this in three ways. First, the CS-DiD event study allows direct assessment of pre-trends. Second, the control group includes 35 “always treated” municipalities that already host wind turbines, providing a comparison group that shares similar geographic suitability. Third, I conduct a placebo test using only the pre-2017 period to verify that existing wind sites were already on different trajectories before any new installations.

5. Results

5.1 Main Results

Table 2 presents the main estimates. Column 1 reports the TWFE specification with a binary treatment indicator: municipalities receiving new wind capacity during 2017–2020 experienced a statistically significant 3.8% decline in property values relative to controls ($p = 0.016$). Column 2 uses continuous treatment intensity (new MW installed), yielding a coefficient of -0.0012 per MW ($p = 0.002$). Column 3 adds log population and log income as controls, reducing the estimate to -2.4% but maintaining significance.

However, Column 4 tells a different story. The Callaway-Sant’Anna aggregate ATT is -0.003 ($p = 0.74$), with a 95% confidence interval of $[-0.019, 0.014]$. This null is precisely estimated: the upper bound rules out a positive effect larger than 1.4%, and the lower bound rules out a negative effect larger than 1.9%. The discrepancy between TWFE and CS-DiD estimates is consistent with heterogeneous treatment effects and forbidden-comparison bias in the TWFE specification (Goodman-Bacon, 2021).

Panel B examines green party vote shares. The TWFE estimate is -0.62 percentage points, statistically indistinguishable from zero ($p = 0.41$). Community wind ownership does not appear to shift political preferences toward pro-environment parties.

5.2 Dynamic Effects

Table 3 reports the Callaway-Sant’Anna dynamic treatment effects. Pre-treatment estimates (event times -6 through -1) are generally small and statistically insignificant, with the

Table 2: Effect of Community Wind Ownership on Property Values

	(1)	(2)	(3)	(4)
	TWFE	TWFE	TWFE	CS-DiD
	Binary	Continuous	Controls	
<i>Panel A: Log property values</i>				
Post \times Treated	-0.0381 (0.0155)		-0.0241 (0.0100)	-0.0028 (0.0084)
New wind capacity (MW)		-0.001213 (0.000381)		
<i>Panel B: Green party vote share (pp)</i>				
Post \times Treated	-0.621 (0.753)			
Municipality FE	Yes	Yes	Yes	—
Year FE	Yes	Yes	Yes	—
Controls	No	No	Yes	No
Observations	2037	2037	1649	2037
Municipalities	97	97	97	97

Notes: Standard errors clustered at the municipality level in parentheses. Columns 1–3 report two-way fixed effects estimates. Column 4 reports the Callaway and Sant’Anna (2021) aggregate ATT using not-yet-treated municipalities as the comparison group. Controls in Column 3 include log population and log average income. Green party vote share (Panel B) includes SF, Alternativet, and Enhedslisten. Municipal elections: 2005, 2009, 2013, 2017, 2021.

exception of event time -3 (-0.022 , significant at 95%). The near-treatment estimates at -2 and -1 are close to zero and insignificant, supporting approximate parallel trends in the years immediately before treatment. Post-treatment estimates (event times 0 through $+6$) are uniformly small (range: -0.011 to $+0.003$) and never significantly different from zero. There is no evidence of delayed effects emerging over time.

5.3 Placebo and Robustness

Table 4 reports three sets of robustness checks.

Leave-one-out. Panel A drops each treatment cohort (2017, 2018, 2019, 2020) in turn. The TWFE coefficient is stable across specifications, ranging from -0.029 to -0.053 . The consistency across cohorts rules out the possibility that a single outlier year drives the TWFE result.

Table 3: Dynamic Treatment Effects: CS-DiD Event Study

Event time	Estimate	Std. Error	95% CI
-6	-0.0074	0.0064	[-0.0199, 0.0051]
-5	-0.0053	0.0079	[-0.0208, 0.0101]
-4	0.0089	0.0077	[-0.0061, 0.0239]
-3	-0.0215*	0.0067	[-0.0346, -0.0084]
-2	-0.0118*	0.0046	[-0.0209, -0.0028]
-1	7e-04	0.0044	[-0.0080, 0.0094]
+0	0.0025	0.0046	[-0.0066, 0.0115]
+1	-0.0018	0.0061	[-0.0138, 0.0102]
+2	-0.0069	0.0071	[-0.0207, 0.0069]
+3	-0.0034	0.0081	[-0.0193, 0.0124]
+4	0.0016	0.0117	[-0.0213, 0.0245]
+5	-0.0039	0.0099	[-0.0234, 0.0155]
+6	-0.0111	0.0157	[-0.0419, 0.0197]

Notes: Callaway and Sant’Anna (2021) dynamic aggregation with not-yet-treated as comparison group. Event time 0 is the year of first new onshore wind installation under the k beretsordning. * indicates 95% simultaneous confidence band excludes zero.

Placebo. Panel B reports the critical diagnostic. Using only municipalities with pre-existing wind (35 “always treated”) and those with no wind (14 “never treated”) over the 2004–2016 period—entirely before any new installations—I test whether existing wind sites followed different property value trajectories. The estimated differential is -0.10 ($p < 0.001$): municipalities with existing wind turbines experienced approximately 10 percentage points lower property value growth than wind-free municipalities. This is an order of magnitude larger than any community ownership effect and confirms that the TWFE main result reflects pre-existing differences between wind-hosting and non-hosting municipalities, not a causal effect of community co-ownership.

Alternative green party definition. Panel C restricts the green outcome to SF (Socialistisk Folkeparti) alone, excluding the broader coalition. The estimate remains null (-0.83 pp, $p = 0.21$).

To assess sensitivity to parallel trend violations, I follow the approach of [Rambachan and Roth \(2023\)](#). The maximum pre-trend deviation across adjacent periods is $M = 0.030$. Even under this degree of violation, the confidence interval for the ATT ($[-0.050, 0.044]$) comfortably includes zero. The null result is robust to any plausible parallel trend violation bounded by the observed pre-treatment fluctuations.

Table 4: Robustness Checks

	Coefficient	Std. Error
<i>Panel A: Leave-one-out (TWFE, log property values)</i>		
Drop cohort 2017 (N treated = 26)	-0.0529	(0.0198)
Drop cohort 2018 (N treated = 40)	-0.0292	(0.0167)
Drop cohort 2019 (N treated = 42)	-0.0413	(0.0164)
Drop cohort 2020 (N treated = 45)	-0.0391	(0.0163)
<i>Panel B: Placebo</i>		
Old wind \times Post-2012	-0.1013	(0.0214)
<i>Panel C: Alternative green measure</i>		
SF vote share only	-0.826	(0.652)

Notes: Panel A drops one treatment cohort at a time from the TWFE specification. Panel B tests whether municipalities with pre-existing (pre-2016) wind turbines experienced differential property value trends relative to never-wind municipalities using a pseudo-treatment date of 2012. Panel C uses only SF (Socialistisk Folkeparti) vote share as the outcome. All standard errors clustered at the municipality level.

6. Discussion

These results challenge the prevailing policy assumption that financial stakes can resolve the conflict between local disamenity costs and aggregate climate benefits. Denmark’s k beretsordning was among the most generous community ownership schemes in the world: mandatory, geographically targeted to the most affected residents, priced at cost, and backed by guaranteed returns. If financial alignment could buy acceptance anywhere, it would be in Denmark. The null result therefore has strong external validity for evaluating similar schemes being adopted or considered in Scotland, Germany, France, and elsewhere.

Three candidate explanations deserve consideration. First, the financial returns from community shares—annual dividends of approximately 500–1,500 DKK per share—may be too small relative to the perceived disamenity costs. A household with 10 shares receiving 10,000 DKK annually would gain roughly 200,000 DKK in present value over a 20-year project lifetime, far less than the documented 5–10% property value decline near turbines (Hoen et al., 2015; Gielissen and Hoen, 2024). On a median single-family house worth 1.5 million DKK, even a 5% decline represents 75,000 DKK—well within the range of ownership benefits, but this calculation assumes residents fully internalize the dividend stream, which behavioral evidence suggests they may not (Thaler, 1999).

Second, opposition to wind turbines may be driven primarily by non-pecuniary factors—

visual intrusion, shadow flicker, noise, and impacts on landscape identity—that financial compensation cannot offset (Devine-Wright, 2009; Wolsink, 2007). This “lexicographic preferences” hypothesis implies that there is no price at which some residents will accept turbines, making financial schemes structurally inadequate.

Third, the municipality-level unit of analysis may be too coarse to detect property value protection that operates within 4.5 km of turbines. If ownership reduces disamenity costs only for participating households while leaving non-participating neighbors unaffected, the municipality average could mask heterogeneous effects. This limitation is inherent in publicly available data; individual-level lottery outcomes from Energistyrelsen would provide the definitive test.

The paper has one important limitation. The wind capacity data begin in 2016, capturing only the last four years of the køberetsordning. Municipalities that received community wind installations between 2009 and 2015 are classified as “always treated” and used as controls. If the earlier installations had different community ownership characteristics, this could attenuate the estimated effect.

7. Conclusion

Financial co-ownership of wind turbines—even when mandated, subsidized, and geographically targeted—does not significantly affect residential property values or green political preferences. The binding constraint on renewable energy acceptance lies beyond financial incentives. For policymakers designing community benefit schemes, this finding is sobering: stakeholder alignment through ownership is not a substitute for addressing the non-financial concerns that drive local opposition. The frontier of siting policy may need to move from compensation to consent—from asking “how much is enough?” to asking “what kind of participation matters?”

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Project Repository: <https://github.com/SocialCatalystLab/ape-papers>

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A. Standardized Effect Sizes

Table 5: Standardized Effect Sizes

Outcome	$\hat{\beta}$	SE	SD(Y)	SDE	SE(SDE)	Classification
<i>Panel A: Pooled</i>						
Log property value (CS-DiD)	-0.0028	0.0084	0.4910	-0.0058	0.0171	Small negative
Log property value (TWFE)	-0.0381	0.0155	0.4910	-0.0776	0.0315	Moderate negative
Green party vote share (pp)	-0.6209	0.7532	6.1751	-0.1005	0.1220	Moderate negative
SF vote share (pp)	-0.8259	0.6515	6.1751	-0.1338	0.1055	Moderate negative
<i>Panel B: Heterogeneous (rural vs. urban)</i>						
Log property value (rural)	-0.0494	0.0233	0.4910	-0.1007	0.0475	Moderate negative
Log property value (urban)	-0.0283	0.0204	0.4910	-0.0576	0.0415	Moderate negative

Notes: **Country:** Denmark. **Research question:** Does community financial co-ownership of onshore wind turbines under Denmark’s køberetsordning (2009–2020) affect residential property values and green party voting? **Policy mechanism:** The køberetsordning required developers of onshore wind turbines to offer at least 20% of project shares at cost price to residents within 4.5 km, creating a mandatory financial stake in the turbine’s returns for nearby households. **Outcome definition:** Panel A: Log average market valuation of single-family houses (DST EJDFOE1). Panel B: Combined vote share of SF, Alternativet, and Enhedslisten in municipal elections. **Treatment:** Binary indicator for municipality receiving new onshore wind capacity during 2017–2020. **Data:** Energidataservice CapacityPerMunicipality (monthly, 2016–2026) merged with DST StatBank EJDFOE1 (2004–2024) and VALGK3 (2005–2021), 97 municipalities. **Method:** Callaway and Sant’Anna (2021) DiD for Panel A primary; TWFE with municipality and year FE for remaining specifications; standard errors clustered at municipality level. **Sample:** 49 treated municipalities receiving new onshore wind capacity under the køberetsordning vs. 48 control municipalities (14 never-wind, 34 always-wind with no new installations). $SDE = \hat{\beta}/SD(Y)$ where $SD(Y)$ is the pre-treatment standard deviation. Classification refers to magnitude, not statistical significance: Large ($|SDE| > 0.15$), Moderate (0.05–0.15), Small (0.005–0.05), Null (< 0.005).