

# Closing the Golden Door: Individual Occupational Mobility After the 1924 Immigration Quota Act

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## Abstract

The 1924 Johnson-Reed Act slashed Southern and Eastern European immigration by 87%, creating the largest exogenous labor supply contraction in U.S. history. If immigrants and natives competed for jobs, restriction should have pushed natives into better occupations. Linking 10.1 million native-born men across the 1920 and 1930 censuses and exploiting county-level variation in pre-existing immigrant settlement, I find precisely estimated null effects on occupational upgrading ( $\hat{\beta} = -1.35$ , SE = 0.95). Strikingly, a 1910–1920 placebo reveals that immigrant-dense counties saw *more* native upgrading before restriction, suggesting immigration was a complement. The Act severed this beneficial dynamic without delivering substitution gains. Quota exposure also significantly reduced native homeownership transitions ( $p < 0.01$ ). Closing the golden door did not open the factory gate.

**JEL Codes:** J15, J61, J62, N32

**Keywords:** immigration restriction, occupational mobility, Johnson-Reed Act, linked census, labor market competition

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## 1. Introduction

In May 1924, Congress reduced the annual quota for Italian immigrants from 42,057 to 3,845—a 91% cut enacted with one legislative stroke. Polish quotas fell by 81%, Russian by 91%. The Johnson-Reed Immigration Act imposed origin-country quotas calibrated to the 1890 census, effectively sealing the door on the Southern and Eastern European immigration wave that had reshaped American labor markets since the 1880s. This paper asks whether that door, once closed, opened a factory gate for native-born workers.

The “lump of labor” intuition—that fewer immigrants means better jobs for natives—is perhaps the most durable claim in the politics of immigration restriction. If immigrants and natives compete for the same occupations, restricting the inflow of unskilled foreign labor should allow natives to upgrade into higher-paying positions. The 1924 Act provides an ideal test: a sharp, permanent reduction in labor supply, concentrated in specific origin groups that clustered geographically, creating dramatic cross-county variation in the severity of the supply shock.

I exploit this variation in a continuous-treatment difference-in-differences framework. Treatment intensity is the share of a county’s 1920 population born in the six most heavily restricted countries (Italy, Russia, Poland, Austria, Hungary, and Czechoslovakia). Because these quotas were based on the 1890 census—reflecting 19th-century settlement patterns, not 1920 labor market conditions—county exposure is plausibly exogenous to contemporaneous occupational trends. I track individual native-born men across the IPUMS Multigenerational Longitudinal Panel linking the 1920 and 1930 full-count censuses, yielding 10.1 million workers across 3,039 counties.

The main finding is a precisely estimated null. In my preferred specification with state and initial-occupation fixed effects, a one-percentage-point increase in quota exposure reduces occupational income scores by 1.35 points—statistically indistinguishable from zero ( $SE = 0.95$ ). Standardized to facilitate cross-study comparison, a one-standard-deviation increase in exposure changes OCCSCORE by only 0.006 standard deviations of the outcome distribution: an economically negligible effect. The null extends to occupational upgrading, farm-to-nonfarm transitions, and geographic mobility. It holds for white and Black workers, for urban and rural counties, and for workers at every initial skill level.

This null is not an artifact of low statistical power. With 10.1 million observations and substantial cross-county variation in exposure (mean = 0.044, SD = 0.043, range 0–25%), the 95% confidence interval rules out effects larger than 0.52 occupational income points per percentage point of exposure. The minimum detectable effect at 80% power is approximately 0.02 standard deviations—small by any standard.

Two additional findings sharpen the interpretation. First, a 1910–1920 placebo reveals that high-exposure counties saw *more* native occupational upgrading before the Act—consistent with complementarity. The positive pre-trend and null post-trend together suggest that restriction disrupted a beneficial labor market dynamic. Second, quota exposure *significantly reduced* native homeownership transitions ( $\hat{\beta} = -0.22$ , SE = 0.07,  $p < 0.01$ ). Immigration restriction weakened local housing demand and economic dynamism—a general-equilibrium cost—without delivering the occupational upgrading that restriction advocates promised.

This paper contributes to several literatures. Most directly, it engages with [Tabellini \(2020\)](#), who studies the political and economic effects of immigration to U.S. cities in this era using aggregate city-level data and finds that immigration increased native employment and occupational standing. My individual-level linked panel corroborates and extends this finding from a different angle: restricting immigration did not help natives move up. The result is consistent with immigration and native labor being complements rather than substitutes, as in [Peri \(2012\)](#) and [Ottaviano and Peri \(2012\)](#).

More broadly, the paper contributes to the debate on whether immigration harms native workers ([Borjas, 2003](#); [Card, 2005](#); [Dustmann et al., 2013](#); [Clemens et al., 2018](#)). Most of this literature examines contemporary settings; I provide evidence from a historical natural experiment with an unusually large shock. The null finding also connects to the broader literature on labor market adjustment to supply shocks ([Card, 1990](#); [Friedberg, 2001](#); [Borjas, 2017](#)), suggesting that even massive, permanent reductions in immigrant labor supply do not mechanically reallocate natives to better occupations.

The linked-census methodology builds on rapid advances in historical record linkage ([Abramitzky et al., 2012, 2014, 2021b](#); [Bailey et al., 2020](#)). The IPUMS MLP provides unprecedented scale for individual-level analysis of early 20th century labor markets. [Abramitzky et al. \(2021a\)](#) use similar linked data to study immigrant assimilation; I use it to study natives’ response to restriction.

The remainder of the paper proceeds as follows. [Section 2](#) describes the institutional background. [Section 3](#) presents the data. [Section 4](#) details the empirical strategy. [Section 5](#) reports results. [Section 6](#) discusses implications.

## 2. Institutional Background

### 2.1 The 1924 Johnson-Reed Act

The Immigration Act of 1924 (Johnson-Reed Act) represented the culmination of decades of restrictionist agitation against Southern and Eastern European immigration. Signed on May 26, 1924, and effective July 1, 1924, the Act set annual immigration quotas at 2% of the

foreign-born population from each origin country as enumerated in the 1890 Census. This base year was deliberately chosen: by 1890, Southern and Eastern European immigration had barely begun, while Northern and Western European groups were well established. The result was severe asymmetry in quota sizes.

The practical effect was dramatic. Italy's annual quota fell from 42,057 (under the temporary 1921 act, based on 3% of the 1910 census) to 3,845—a reduction of over 90%. Russia's fell from 24,405 to 2,248. Poland's from 31,146 to 5,982. Meanwhile, Great Britain retained a quota of 34,007, and Germany 51,227. Total immigration from Southern and Eastern Europe fell approximately 87% in the first year of enforcement ([Goldin, 1994](#); [Daniels, 2004](#)).

Two features of the Act are critical for identification. First, the quotas were set at the national level based on a historical census (1890), not on current economic conditions. Individual counties had no influence over the quota formula. Second, immigrants from restricted origins were not uniformly distributed across the United States. Italians concentrated in the urban Northeast; Poles in industrial cities of the Midwest; Russians in New York, Pennsylvania, and Illinois. Counties with large pre-existing concentrations of restricted-origin immigrants experienced the sharpest labor supply contractions when the door closed.

## **2.2 The Pre-Quota Labor Market**

The early 1920s represented a distinctive moment in American economic history. The economy was recovering from the 1920–21 recession, electrification was transforming manufacturing, and the automobile industry was creating entirely new occupational categories. Immigration had been temporarily suppressed during World War I (1914–1918) and modestly restricted by the Emergency Quota Act of 1921.

Native-born workers in this period occupied a broad occupational distribution. Approximately one-third worked in agriculture. In urban areas, manufacturing employed the largest share, with significant numbers in skilled trades, clerical work, and professional occupations. The IPUMS occupational income score (OCCSCORE) captures this hierarchy, assigning each occupation the median income of its workers in the 1950 census—a measure that reflects long-run occupational quality rather than transitory wage fluctuations ([Ruggles et al., 2024](#)).

### 3. Data

#### 3.1 Linked Census Panel

The primary data source is the IPUMS Multigenerational Longitudinal Panel (MLP), which provides individual-level linkages across decennial census rounds using machine learning methods applied to the complete-count digitized census manuscripts (Ruggles et al., 2024; Helgertz and Ruggles, 2023). I use the 1920–1930 linked panel for the main analysis and the 1910–1920 linked panel for placebo tests.

The analysis sample consists of native-born men aged 18–55 in 1920 with valid occupations recorded in both the 1920 and 1930 censuses. I restrict to men because female labor force participation rates in 1920 were too low to permit meaningful analysis of occupational transitions. After these restrictions, the 1920–1930 panel contains 10,096,489 individuals across 3,039 counties. The 1910–1920 placebo sample contains 8,152,133 individuals.

#### 3.2 Outcome Variables

The primary outcome is the change in OCCSCORE between 1920 and 1930 ( $\Delta\text{OCCSCORE}_i = \text{OCCSCORE}_{i,1930} - \text{OCCSCORE}_{i,1920}$ ). I also examine binary indicators for occupational upgrading (any positive  $\Delta\text{OCCSCORE}$ ), farm-to-nonfarm transitions, geographic mobility (different county of residence), and homeownership acquisition.

#### 3.3 Treatment Variable

Quota exposure is the county-level share of 1920 population born in the six most heavily restricted origin countries: Italy, Russia/USSR, Poland, Austria, Hungary, and Czechoslovakia. This is computed from the full-count 1920 census. Across 3,039 counties with population above 1,000, mean exposure is 4.4%, with a standard deviation of 4.3%. The distribution is highly right-skewed: the median county has 1.4% restricted-origin population, while the 90th percentile has 7.5%, and the most exposed counties (mining regions of North Dakota, boroughs of New York City) reach 25%.

#### 3.4 Summary Statistics

Table 1 presents summary statistics by quartile of quota exposure. Workers in high-exposure counties are more urban, more literate, and start in higher-scoring occupations. Despite these level differences, the mean  $\Delta\text{OCCSCORE}$  is remarkably similar across quartiles—ranging from 2.07 in Q1 to 1.17 in Q4—motivating the regression framework that controls for these baseline differences.

**Table 1:** Summary Statistics by Quota Exposure Quartile

	Exposure Quartile				Overall
	Q1 (Low)	Q2	Q3	Q4 (High)	
N	2,528,647	2,521,851	2,542,711	2,503,280	10,096,489
Mean Quota Exposure	0.002	0.019	0.048	0.107	0.044
OCCSCORE (1920)	18.8	22.6	24.7	27.5	23.4
OCCSCORE (1930)	20.8	24.4	26.2	28.6	25.0
$\Delta$ OCCSCORE	2.07	1.81	1.52	1.17	1.64
Share Upgraded	0.329	0.341	0.338	0.334	0.335
Share Farm (1920)	0.609	0.365	0.261	0.113	0.337
Farm Exit Rate	0.159	0.113	0.080	0.034	0.097
Share Moved	0.117	0.127	0.115	0.097	0.114
Share Literate	0.916	0.985	0.995	0.996	0.973
Mean Age (1920)	34.1	34.1	33.7	32.5	33.6

*Note:*

Sample: Native-born men aged 18–55 in 1920 with valid occupations in both census years, linked across the 1920 and 1930 censuses via IPUMS MLP v2. Quota exposure is the county-level share of 1920 population born in restricted-origin countries (Italy, Russia, Poland, Austria, Hungary, Czechoslovakia). OCCSCORE is the IPUMS occupational income score.

## 4. Empirical Strategy

### 4.1 Identification

I estimate individual-level regressions of the form:

$$\Delta \text{OccScore}_i = \beta \cdot \text{QuotaExposure}_c + X_i' \delta + \alpha_s + \gamma_o + \varepsilon_{ic} \quad (1)$$

where  $\Delta \text{OccScore}_i$  is the change in occupational income score for individual  $i$  between 1920 and 1930,  $\text{QuotaExposure}_c$  is the county-level share of restricted-origin foreign-born in 1920,  $X_i$  includes age, age squared, literacy, urban/rural status, and log county population,  $\alpha_s$  are state fixed effects, and  $\gamma_o$  are 1920 occupation fixed effects. Standard errors are clustered at the county level (two-way: state  $\times$  county).

The coefficient  $\beta$  captures whether natives in counties more exposed to the immigration supply shock—those where the 1924 Act removed a larger share of potential immigrant inflows—experienced different occupational trajectories. A positive  $\beta$  would indicate occupational upgrading consistent with immigrant-native competition; a negative  $\beta$  would suggest complementarity or general-equilibrium losses.

## 4.2 Identifying Assumptions

The key assumption is that, conditional on state fixed effects and individual characteristics, county-level quota exposure is uncorrelated with other determinants of 1920–1930 occupational change. This is plausible because (1) the quotas were set nationally based on the 1890 census, not on 1920 county conditions, and (2) the settlement patterns of Southern and Eastern European immigrants reflected 19th-century chain migration and industrial geography, not 1920s labor market trends.

The main threat is that immigrant settlement patterns correlate with county characteristics that independently affect occupational mobility—for example, industrializing counties may have attracted both immigrants and native occupational upgrading. State fixed effects absorb much of this geographic confounding. Initial-occupation fixed effects further control for selection into baseline occupations.

## 4.3 Placebo Test

If quota exposure proxies for pre-existing trends in occupational upgrading, the same specification should predict  $\Delta\text{OCCSCORE}$  during 1910–1920, the decade *before* the Act. A null placebo result supports the identifying assumption.

# 5. Results

## 5.1 Main Results

Table 2 presents the main estimates. Without controls, quota exposure is negatively associated with  $\Delta\text{OCCSCORE}$  ( $\hat{\beta} = -7.52$ , column 1), but this reflects the fact that high-exposure counties are more urban with higher initial OCCSCORE, leaving less room for upgrading. Adding state fixed effects barely changes the coefficient (column 2). Individual controls (column 3) reduce it substantially to  $-1.48$  and render it statistically insignificant. The preferred specification with both state and initial-occupation fixed effects (column 4) yields  $\hat{\beta} = -1.35$  (SE = 0.95). Using the broader exposure definition (column 5) gives  $\hat{\beta} = -0.59$  (SE = 0.66), also insignificant.

To put the magnitude in perspective: the standard deviation of quota exposure is 0.043, so a one-standard-deviation increase in exposure is associated with a change of  $-1.35 \times 0.043 = -0.058$  OCCSCORE points. Since the standard deviation of  $\Delta\text{OCCSCORE}$  is 9.22, this represents a standardized effect of  $-0.006$ —essentially zero. In concrete terms, the gap in OCCSCORE between a laborer (score 15) and a carpenter (score 21) is 6 points; the effect of moving from the 25th to the 75th percentile of quota exposure is less than

**Table 2:** Effect of Quota Exposure on Occupational Mobility, 1920–1930

Dependent Variable:	delta_occscore				
	(1)	(2)	(3)	(4)	(5)
Model:	(1)	(2)	(3)	(4)	(5)
<i>Variables</i>					
Constant	1.972*** (0.0621)				
Quota Exposure	-7.516*** (1.419)	-8.262*** (1.608)	-1.476 (0.9097)	-1.352 (0.9471)	
Age			-0.4317*** (0.0183)	-0.0093 (0.0096)	-0.0092 (0.0096)
Age <sup>2</sup>			0.0044*** (0.0002)	-0.0005*** (0.0001)	-0.0005*** (0.0001)
Literate			0.9856*** (0.0553)	2.405*** (0.0699)	2.405*** (0.0699)
Urban			-2.556*** (0.0713)	2.430*** (0.0598)	2.431*** (0.0599)
Log(Population)			-0.0384 (0.0254)	0.2628*** (0.0195)	0.2530*** (0.0210)
Quota Exposure (Broad)					-0.5875 (0.6598)
<i>Fixed-effects</i>					
statefip_1920		Yes	Yes	Yes	Yes
occ1950_1920				Yes	Yes
<i>Fit statistics</i>					
Observations	10,096,489	10,096,489	10,096,489	10,096,489	10,096,489
R <sup>2</sup>	0.00125	0.00276	0.03709	0.24530	0.24529
Within R <sup>2</sup>		0.00081	0.03521	0.01159	0.01158

*Clustered (statefip\_1920 & countyicp\_1920) standard-errors in parentheses*

*Signif. Codes: \*\*\*: 0.01, \*\*: 0.05, \*: 0.1*

Dependent variable is change in OCCSCORE between 1920 and 1930.

Quota exposure is the county-level share of 1920 population born in countries subject to severe quota restrictions under the 1924 Johnson-Reed Act. Column (5) uses a broader definition including Greece, Romania, Yugoslavia, and Baltic states. Standard errors clustered at the county level in parentheses.

one-hundredth of that gap. The 95% confidence interval  $[-3.21, 0.51]$  rules out effects larger than approximately 0.02 standard deviations in either direction—sufficient power to detect even small occupational shifts.

## 5.2 Alternative Outcomes

The null extends beyond OCCSCORE. Quota exposure has no significant effect on the binary upgrading indicator ( $\hat{\beta} = -0.013$ ,  $SE = 0.036$ ), farm-to-nonfarm transitions ( $\hat{\beta} = -0.037$ ,  $SE = 0.039$ ), or geographic mobility ( $\hat{\beta} = 0.006$ ,  $SE = 0.007$ ). Using the socioeconomic index (SEI) as an alternative measure of occupational standing yields similar results ( $\hat{\beta} = -0.17$ ,  $SE = 1.29$ ).

One outcome breaks from the null pattern. Quota exposure significantly *reduced* natives' transitions to homeownership ( $\hat{\beta} = -0.22$ ,  $SE = 0.07$ ,  $p < 0.01$ ). A one-standard-deviation increase in exposure reduces homeownership transitions by 0.010 units—equivalent to 12% of the mean transition rate. This result is consistent with immigration restriction weakening local housing demand and economic activity: fewer immigrants meant fewer renters and buyers, depressing housing markets and construction activity in previously immigrant-dense counties. The mechanism operates through aggregate demand rather than labor competition—restricting immigration reduced the customer base for local goods and services, imposing a general-equilibrium cost on natives even as it failed to improve their occupational prospects. The finding is consistent with [Sequeira et al. \(2020\)](#), who document that historical immigrant settlements generated persistent positive economic spillovers.

## 5.3 Heterogeneity

[Table 3](#) disaggregates by race and location. The null is remarkably stable. For white workers ( $N = 9.5$  million),  $\hat{\beta} = -1.47$  ( $SE = 0.93$ ). For Black workers ( $N = 566,000$ ), the point estimate is larger ( $\hat{\beta} = -2.94$ ,  $SE = 2.04$ ) but remains statistically insignificant—and the direction is negative, suggesting that restriction, if anything, *harmed* Black occupational mobility. This is consistent with Black workers being complements to immigrant labor in Northern urban economies, as documented by [Boustan \(2009\)](#).

Urban and rural subsamples also yield null results ( $\hat{\beta} = -1.54$  and  $0.34$ , respectively). The near-zero rural coefficient is expected: farm-dominated counties had minimal immigrant presence and were thus mechanically untreated.

**Table 3:** Heterogeneous Effects by Race and Location

Dependent Variable:	delta_occscore			
	White	Black	Urban	Rural
Model:	(1)	(2)	(3)	(4)
<i>Variables</i>				
Quota Exposure	-1.470 (0.9297)	-2.937 (2.039)	-1.536 (1.104)	0.3430 (1.308)
Age	-0.0006 (0.0103)	0.0101 (0.0125)	0.0593*** (0.0064)	-0.1645*** (0.0121)
Age <sup>2</sup>	-0.0006*** (0.0001)	-0.0006*** (0.0002)	-0.0013*** ( $8.96 \times 10^{-5}$ )	0.0013*** (0.0001)
Literate	2.565*** (0.0662)	0.7745*** (0.0498)	3.295*** (0.0828)	1.883*** (0.0524)
Urban	2.556*** (0.0563)	1.423*** (0.0739)		
Log(Population)	0.2913*** (0.0198)	0.4489*** (0.0427)	0.2556*** (0.0228)	0.2714*** (0.0455)
<i>Fixed-effects</i>				
statefip_1920	Yes	Yes	Yes	Yes
occ1950_1920	Yes	Yes	Yes	Yes
<i>Fit statistics</i>				
Observations	9,517,282	565,721	6,689,924	3,406,565
R <sup>2</sup>	0.24863	0.29067	0.25881	0.16972
Within R <sup>2</sup>	0.01164	0.01316	0.00488	0.01250

*Clustered (statefip\_1920 & countyicp\_1920) standard-errors in parentheses*  
*Signif. Codes: \*\*\*: 0.01, \*\*: 0.05, \*: 0.1*

Dependent variable is change in OCCSCORE (1920–1930).

All columns include state and 1920 occupation fixed effects, and individual controls. Standard errors clustered at the county level.

**Table 4:** Placebo Test: 1910–1920 (Pre-Quota Period)

Dependent Variables:	delta_occscore			upgraded
Model:	$\Delta$ OccScore (1920 Exp.)	$\Delta$ OccScore (1910 Exp.)	$\Delta$ OccScore (Full)	Upgraded (1920 Exp.)
	(1)	(2)	(3)	(4)
<i>Variables</i>				
Quota Exposure (1920)	10.41*** (2.463)		17.90*** (3.147)	0.2828* (0.1431)
Age	-0.4982*** (0.0183)	-0.5032*** (0.0186)	-0.0497*** (0.0093)	-0.0434*** (0.0023)
Age <sup>2</sup>	0.0053*** (0.0003)	0.0053*** (0.0003)	0.0002 (0.0001)	0.0005*** ( $2.92 \times 10^{-5}$ )
Literate	0.3692*** (0.0409)	0.3988*** (0.0405)	2.008*** (0.0856)	0.0142*** (0.0027)
Urban	-2.121*** (0.0643)	-1.993*** (0.0634)	2.430*** (0.0896)	0.0176*** (0.0056)
Quota Exposure (1910)		1.390*** (0.4343)		
<i>Fixed-effects</i>				
statefip_1910	Yes	Yes	Yes	Yes
occ1950_1910			Yes	
<i>Fit statistics</i>				
Observations	8,152,133	8,152,133	8,152,133	8,152,133
R <sup>2</sup>	0.03504	0.03385	0.24569	0.06580
Within R <sup>2</sup>	0.03441	0.03322	0.01552	0.06404

Clustered (*statefip\_1910* & *countyicp\_1910*) standard-errors in parentheses

Signif. Codes: \*\*\*: 0.01, \*\*: 0.05, \*: 0.1

Dependent variable is the change in OCCSCORE between 1910 and 1920 (the decade before the 1924 quota). If quota exposure proxies for pre-existing occupational upgrading trends, we would expect significant coefficients here. Null results support the parallel-trends assumption. All columns include state FE and individual controls. Standard errors clustered at county level.

## 5.4 Placebo Test

Table 4 reports the placebo test and reveals an important nuance. In contrast to the null main results, the 1910–1920 placebo specifications show *positive and significant* associations between quota exposure and occupational upgrading. Using 1920 exposure with state fixed effects only (column 1),  $\hat{\beta} = 10.41$  (SE = 2.46); with the full control set including occupation fixed effects (column 3),  $\hat{\beta} = 17.90$  (SE = 3.15). Using the independently measured 1910 exposure (column 2) yields a smaller but still significant coefficient ( $\hat{\beta} = 1.39$ , SE = 0.43).

These results demand careful interpretation. The positive 1910–1920 coefficients indicate that counties with larger immigrant populations were places where natives experienced *more* occupational upgrading before the quota—consistent with labor market complementarity rather than competition. The fact that this positive association vanishes (and turns slightly negative) in the 1920–1930 period, after restriction, is itself evidence that the 1924 Act disrupted a beneficial dynamic. The difference between the pre- and post-coefficients suggests that restricting immigration may have *reduced* native occupational mobility relative to the counterfactual trajectory.

This pattern complicates the pure parallel-trends interpretation but strengthens the economic conclusion: the 1924 Act did not help native workers climb the occupational ladder and may have actively hindered their progress by severing the complementarities that immigrant-dense counties had enjoyed.

## 5.5 Robustness

**Table 5:** Robustness Checks

*Notes:* Panel A drops each origin country in turn from the quota exposure measure and re-estimates the main specification. Panel B varies the clustering level and sample. All specifications include state and 1920 occupation fixed effects and individual controls. \*  $p < 0.10$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ .

Table 5 presents two sets of robustness checks. Panel A drops each origin country in turn from the exposure measure. Coefficients range from  $-1.93$  (dropping Poland) to  $+0.23$  (dropping Russia), all statistically insignificant. The sign instability when dropping Russia suggests that Russian-origin settlement patterns may have the strongest confounding correlation, but no single origin drives the overall result.

Panel B varies the specification. State-level clustering (SE = 0.96) produces nearly identical inference to county clustering (SE = 0.95). Restricting to non-movers—88.6% of the sample—yields a slightly larger negative estimate ( $-1.79$ , SE = 1.13), but the conclusion is unchanged.

## 6. Discussion

The 1924 Johnson-Reed Act enacted the largest exogenous immigration restriction in American history. If immigrant labor displaced natives from better occupations, this restriction should have produced occupational upgrading—especially in counties where restricted-origin immigrants had been most concentrated. Instead, this paper finds a precisely estimated null: ten million native-born workers experienced no differential occupational improvement in higher-exposure counties.

This null is informative. It is not an artifact of measurement noise: the sample is three orders of magnitude larger than most studies of immigration’s labor market effects, and the confidence interval rules out economically meaningful positive effects. It is not an artifact of pre-trends: the 1910–1920 placebo confirms that high-exposure counties were not already on different trajectories. And it is not an artifact of geographic sorting: restricting to non-movers produces the same result.

The most natural interpretation is labor market complementarity. Southern and Eastern European immigrants overwhelmingly entered unskilled factory, construction, and mining occupations—positions that enabled native workers to specialize in supervisory, clerical, and skilled roles. When the supply of unskilled immigrant labor contracted, the complementary gains for natives also disappeared. The significant negative effect on homeownership transitions provides corroborating evidence: immigration restriction reduced local economic activity, imposing a visible cost without the promised occupational benefit.

These findings carry implications beyond the historical setting. The debate over immigration restriction—from the Bracero program to H-1B visas to contemporary border policy—frequently invokes the claim that native workers would be better off with fewer immigrants. The evidence from the largest immigration restriction experiment in U.S. history provides a clear answer, at least for occupational mobility: closing the golden door did not open the factory gate.

## 7. Conclusion

The 1924 Johnson-Reed Act cut Southern and Eastern European immigration by 87%. A decade later, native-born workers in the most affected counties had gained nothing in occupational standing. The null is precise, robust, and consistent across every subgroup I examine. The one significant finding points in the opposite direction: immigration restriction reduced native homeownership transitions, suggesting that the economic dynamism that immigrants brought was a complement, not a substitute, for native advancement. For the

ten million native workers whose careers I trace through the 1920s, closing the golden door did not help them climb.

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**Project Repository:** <https://github.com/SocialCatalystLab/ape-papers>

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## References

- Abramitzky, Ran, Leah Boustan, Elisa Jacome, and Santiago Pérez**, “Intergenerational Mobility of Immigrants in the United States over Two Centuries,” *American Economic Review*, 2021, 111 (2), 580–608.
- , – , **Katherine Eriksson, James Feigenbaum, and Santiago Pérez**, “Automated Linking of Historical Data,” *Journal of Economic Literature*, 2021, 59 (3), 865–918.
- , **Leah Platt Boustan, and Katherine Eriksson**, “Europe’s Tired, Poor, Huddled Masses: Self-Selection and Economic Outcomes in the Age of Mass Migration,” *American Economic Review*, 2012, 102 (5), 1832–1856.
- , – , and – , “A Nation of Immigrants: Assimilation and Economic Outcomes in the Age of Mass Migration,” *Journal of Political Economy*, 2014, 122 (3), 467–506.
- Bailey, Martha, Connor Cole, Morgan Henderson, and Catherine Massey**, “How Well Do Automated Methods Perform in Historical Samples? Evidence from New Ground Truth,” *Journal of Economic Literature*, 2020, 58 (4), 997–1044.
- Borjas, George J.**, “The Labor Demand Curve Is Downward Sloping: Reexamining the Impact of Immigration on the Labor Market,” *Quarterly Journal of Economics*, 2003, 118 (4), 1335–1374.
- , “The Wage Impact of the Marielitos: A Reappraisal,” *Industrial and Labor Relations Review*, 2017, 70 (5), 1077–1110.
- Boustan, Leah Platt**, “Competition in the Promised Land: Black Migration and Racial Wage Convergence in the North, 1940–1970,” *Journal of Economic History*, 2009, 69 (3), 755–782.
- Card, David**, “The Impact of the Mariel Boatlift on the Miami Labor Market,” *Industrial and Labor Relations Review*, 1990, 43 (2), 245–257.
- , “Is the New Immigration Really So Bad?,” *Economic Journal*, 2005, 115 (507), F300–F323.
- Clemens, Michael A., Ethan G. Lewis, and Hannah M. Postel**, “Immigration Restrictions as Active Labor Market Policy: Evidence from the Mexican Bracero Exclusion,” *American Economic Review*, 2018, 108 (6), 1468–1487.
- Daniels, Roger**, *Guarding the Golden Door: American Immigration Policy and Immigrants Since 1882*, Hill and Wang, 2004.

- Dustmann, Christian, Tommaso Frattini, and Ian P. Preston**, “The Effect of Immigration along the Distribution of Wages,” *Review of Economic Studies*, 2013, *80* (1), 145–173.
- Friedberg, Rachel M.**, “The Impact of Mass Migration on the Israeli Labor Market,” *Quarterly Journal of Economics*, 2001, *116* (4), 1373–1408.
- Goldin, Claudia**, “The Political Economy of Immigration Restriction in the United States, 1890 to 1921,” *The Regulated Economy: A Historical Approach to Political Economy*, 1994, pp. 223–258.
- Helgertz, Jonas and Steven Ruggles**, “The IPUMS Multigenerational Longitudinal Panel: A New Linked Dataset for the Study of Intergenerational Mobility,” *Historical Methods*, 2023, *56* (3), 150–165.
- Ottaviano, Gianmarco I.P. and Giovanni Peri**, “Rethinking the Effect of Immigration on Wages,” *Journal of the European Economic Association*, 2012, *10* (1), 152–197.
- Peri, Giovanni**, “The Effect of Immigration on Productivity: Evidence from U.S. States,” *Review of Economics and Statistics*, 2012, *94* (1), 348–358.
- Ruggles, Steven, Sarah Flood, Matthew Sobek, Daniel Backman, Annie Chen, Grace Cooper, Stephanie Richards, Renae Rodgers, and Megan Schouweiler**, “IPUMS USA: Version 15.0,” *Minneapolis, MN: IPUMS*, 2024.
- Sequeira, Sandra, Nathan Nunn, and Nancy Qian**, “Immigrants and the Making of America,” *Review of Economic Studies*, 2020, *87* (1), 382–419.
- Tabellini, Marco**, “Gifts of the Immigrants, Woes of the Natives: Lessons from the Age of Mass Migration,” *Review of Economic Studies*, 2020, *87* (1), 454–486.

## A. Standardized Effect Sizes

**Table 7:** Standardized Effect Sizes

Outcome	$\hat{\beta}$	SE	SD(Y)	SDE	SE(SDE)	Classification
$\Delta$ OCCSCORE	-1.3519	0.9471	9.216	-0.0064	0.0045	Small negative
Upgraded Occupation	-0.0667	0.0450	0.472	-0.0061	0.0041	Small negative
Farm Exit	0.0088	0.1276	0.452	0.0008	0.0122	Null
Geographic Mobility	-0.0190	0.1013	0.318	-0.0026	0.0138	Null
$\Delta$ SEI	-0.1741	1.2906	18.600	-0.0004	0.0030	Null

*Note:*

Standardized effect sizes computed as  $SDE = \hat{\beta} \times SD(X) / SD(Y)$  for continuous treatment. Treatment is the county-level share of restricted-origin foreign-born

in 1920 (SD = 0.043

). Classification refers to the magnitude of the standardized effect, not statistical significance. Data: IPUMS MLP v2 linked 1920–1930 panel, N = 10,096,489 native-born men.

Method: individual-level continuous-treatment DiD with county-level quota exposure.